# Weak Convergence of a Metropolis Algorithm for Bimodal Target Distributions

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#### 1 Random Walk Metropolis algorithm

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## 1 Random Walk Metropolis algorithm

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Summary and Future work

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# MCMC in a nutshell

- Probability measure Π
- Want to estimate (finite) expectation  $\int h \, d\Pi$
- Idea: create a Markov chain  $X_1, X_2, ...$  that has target distribution  $\Pi$
- If the chain is well behaved (irreducible, aperiodic, Harris recurrent),

$$\frac{1}{n}\sum_{i=1}^n h(X_i) \longrightarrow \int h\,d\Pi \quad \text{a.s.}$$

 In general, less efficient than iid sampling, but iid sampling is sometimes unrealizable

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## The Random Walk Metropolis algorithm

- RWM is a very simple MCMC algorithm
- $\pi$  is a (possibly unnormalized) density on  $\mathbb{R}^d$
- Initialize X<sub>0</sub>
- Given X<sub>t</sub>, independently generate Y symmetric around 0
- Set

$$\mathbf{X}_{t+1} = \begin{cases} \mathbf{X}_t + \mathbf{Y}, & \text{w.p.} \quad \alpha_t \\ \mathbf{X}_t, & \text{w.p.} \quad 1 - \alpha_t \end{cases}$$

- $\alpha_t = \min\left\{1, \frac{\pi(\mathbf{X}_t + \mathbf{Y})}{\pi(\mathbf{X}_t)}\right\}$
- Called accept/reject step

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## The Random Walk Metropolis algorithm

Under mild conditions on instrumental distribution (distribution of increments Y),

$$\frac{1}{n}\sum_{t=1}^{n}h(\mathsf{X}_{t})\longrightarrow \frac{\int h(\mathsf{x})\pi(\mathsf{x})\,d\mathsf{x}}{\int \pi(\mathsf{x})\,d\mathsf{x}} \quad \text{a.s.}$$

- Usually  $\mathbf{Y} \sim \mathbf{N}_d \left( \mathbf{0}, \sigma^2 \mathbf{I}_d \right)$
- In theory, works for any  $\sigma > 0...$
- But small  $\sigma \Rightarrow$  Small steps  $\Rightarrow$  Slow exploration
- And large  $\sigma \Rightarrow X_t + Y$  far from the mode  $\Rightarrow \frac{\pi(X_t+Y)}{\pi(X_t)}$  small  $\Rightarrow$  Most steps are rejected  $\Rightarrow$  Slow exploration
- Must choose  $\sigma$  carefully

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# Optimal scaling

- [Roberts et al., 1997] proposed optimal scaling (in simplified framework)
- Assume  $\pi(x) = \prod_{i=1}^d f(x_i)$
- Let  $\sigma = \frac{\ell}{\sqrt{d}}$  (Large dimension  $\Rightarrow$  Small steps)
- Let X<sup>(d)</sup>(t) be Markov chain obtained (with step size depending on dimension d)
- And X<sub>j</sub><sup>(d)</sup>(t) its j-th component
- Smaller steps  $\Rightarrow$  More steps needed, so accelerate chain:  $Z_j^{(d)}(t) = X_j^{(d)}(\lfloor dt \rfloor)$
- Let  $d \to \infty$

#### Random Walk Metropolis algorithm

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## Accelerating the Markov chain



Figure: Trace of the accelerated first component of the Markov chain,  $Z_1^{(d)}$ , for different values of d. The target distribution is multivariate standard Gaussian.

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# Optimal scaling

- Find that  $Z_j^{(d)}$  converges (weakly in Skorokhod topology) to Langevin diffusion with speed measure  $v(\ell) = 2\ell^2 \Phi\left(-\frac{\ell\sqrt{B}}{2}\right)$  and stationnary distribution f
- B is unknown parameter of f
- Find a value  $\hat{\ell}$  that maximizes  $v(\ell)$
- $\hat{\ell}$  is the only value that makes the asymptotic acceptance probability 0.234 (regardless of *B*)!
- So just simulate a few short runs and tune  $\sigma$  so that acceptance rate is roughly 0.234

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## The problem with bimodal distributions

- Suppose the target density  $\pi$  (on  $\mathbb{R}^d$ ) has two distinct modes with a "hole" in between (low density region)
- Optimal scaling strategy tends to favor local exploration (Relatively small  $\sigma$ )
- $\bullet\,$  Small steps  $\Rightarrow\,$  Almost impossible to cross the hole in 1 step
- Almost impossible to accept steps into the hole  $\left(\frac{\pi(X_t+Y)}{\pi(X_t)}\right)$  very small)
- Chain gets stuck in a mode and never explores the other one

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## Solution: A new instrumental distribution

- Previously, Y  $\sim$  N<sub>d</sub>  $\left(0, \sigma^2 I_d\right)$
- Now, let

$$\mathsf{Y} \sim egin{cases} \mathsf{N}_d\left(0,\sigma^2 \mathit{I}_d
ight), & ext{w.p.} & 1-p \ \mathcal{D}, & ext{w.p.} & p \end{cases}$$

where  $\mathcal{D}$  is any distribution on  $\mathbb{R}^d$  symmetric around 0,  $p \in (0,1)$ 

- ullet In practice, choose  $\mathcal D$  to favor switching modes
- Turns the "slowly mixing" chain into a "rapidly mixing" chain ([Guan and Krone, 2007])

### Random Walk Metropolis algorithm

## 2 Weak convergence of our algorithm

- Framework
- The limiting processes



# Framework

- Assume  $\pi(x) = f_1(x_1) \prod_{i=2}^d f(x_i)$
- Instrumental distribution  $\mathcal{D}$ :  $Y_1 \sim \mathcal{D}_1$ ,  $Y_2$ , ...,  $Y_d \sim N(0, \sigma^2)$ , all independent
- Scale both  $\sigma=rac{\ell}{\sqrt{d}}$  and  $p=1\wedgerac{eta}{d},$  but not  $\mathcal{D}_1$
- Accelerate the chain by a factor of *d* (like before)
- The one-dimensional accelerated processes  $Z_i^{(d)}$  weakly converge

# The Langevin diffusion with Metropolis jumps

- $Z_L$  and  $Z_{LM}$  are Langevin diffusions with speed measure  $v(\ell)$  and stationnary distributions f and  $f_1$
- At random times  $T_1, T_2, ...$  generated by a Poisson process of rate  $\beta$ , generate  $Y(T_i) \stackrel{iid}{\sim} \mathcal{D}_1$
- At time  $T_i$ ,  $Z_{LM}$  jumps by  $Y(T_i)$  with probability  $\alpha(\ell, Z_{LM}(T_i^-), Z_{LM}(T_i^-) + Y(T_i))$

$$\alpha(\ell, x, y) = \Phi\left(\frac{\log\frac{f_1(y)}{f_1(x)} - \frac{\ell^2}{2}B}{\ell\sqrt{B}}\right) + \frac{f_1(y)}{f_1(x)}\Phi\left(\frac{-\log\frac{f_1(y)}{f_1(x)} - \frac{\ell^2}{2}B}{\ell\sqrt{B}}\right)$$

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Framework The limiting processes

## The limiting processes

#### Theorem

If the chain starts at stationnarity  $(X^{(d)}(0) \sim \pi)$ , then as  $d \to \infty$ ,

$$Z_1^{(d)} \Rightarrow Z_{LM}, \quad \text{and for } j \ge 2, \quad Z_i^{(d)} \Rightarrow Z_L.$$

Here,  $\Rightarrow$  represents weak convergence in the Skorokhod topology.

Framework The limiting processes

# Optimizing the limiting processes

- Run classic RWM on  $\pi$  and tune  $\sigma$  so that acceptance rate is roughly 0.234
- Run RWM on  $f_1$  with instrumental distribution  $\mathcal{D}_1$  and acceptance probability  $\alpha(\ell, x, y)$  to estimate acceptance probability of large steps, say  $\lambda$
- Choose p, proportion of accepted large steps will be  $pprox p\lambda$
- Number of times we switch modes  $\approx np\lambda$

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- Can work for much more general target distributions (align modes with rotation)
- Find objective way to choose *p* through non-asymptotics
- Can replace the large step distribution  $\mathcal{D}$  by any *algorithm* ("something-inside-Metropolis")



algorithms. Ann. Appl. Probab., 7:110-120.